STA 360/602L: MODULE 7.1

THE METROPOLIS ALGORITHM

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INTRODUCTION

- As a refresher, suppose $y = (y_1, \ldots, y_n)$ and each $y_i \sim p(y|\theta)$. Suppose we specify a prior $\pi(\theta)$ on θ .
- Then as usual, we are interested in

$$\pi(heta|y) = rac{\pi(heta)p(y,| heta)}{p(y)}.$$

- As we already know, it is often difficult to compute p(y).
- Using the Monte Carlo method or Gibbs sampler, we have seen that we don't need to know p(y).
- As long as we have conjugate and semi-conjugate priors, we can generate samples directly from π(θ|y).
- What happens if we cannot sample directly from $\pi(\theta|y)$?



MOTIVATING EXAMPLE

- To motivate our discussions on the Metropolis algorithm, let's explore a simple example.
- Suppose we wish to sample from the following density

$$\pi(heta|y) \propto \exp^{-rac{1}{2} heta^2} + rac{1}{2} \exp^{-rac{1}{2}(heta-3)^2}$$

- This is a mixture of two normal densities, one with mode near 0 and the other with mode near 3.
- Note: we will cover finite mixture models properly soon.
- Anyway, let's use this density to explore the main ideas behind the Metropolis sampler.
- By the way, as you will see, we don't actually need to know the normalizing constant for Metropolis sampling but for this example, find it for practice!

MOTIVATING EXAMPLE

• Let's take a look at the (normalized) density:

 $\pi(\theta|y)$



 There are other ways of sampling from this density, but let's focus specifically on the Metropolis algorithm here.

- From a sampling perspective, we need to have a large group of values, $\theta^{(1)}, \ldots, \theta^{(S)}$ from $\pi(\theta|y)$ whose empirical distribution approximates $\pi(\theta|y)$.
- That means that for any two values a and b, we want

$$\frac{\#\theta^{(s)}=a}{S}\div\frac{\#\theta^{(s)}=b}{S}=\frac{\#\theta^{(s)}=a}{S}\times\frac{S}{\#\theta^{(s)}=b}=\frac{\#\theta^{(s)}=a}{\#\theta^{(s)}=b}\approx\frac{\pi(\theta=a|y)}{\pi(\theta=b|y)}$$

- Basically, we want to make sure that if a and b are plausible values in $\pi(\theta|y)$, the ratio of the number of the $\theta^{(1)}, \ldots, \theta^{(S)}$ values equal to them properly approximates $\frac{\pi(\theta = a|y)}{\pi(\theta = b|y)}$.
- How might we construct a group like this?



- Suppose we have a working group θ⁽¹⁾,...,θ^(s) at iteration s, and need to add a new value θ^(s+1).
- Consider a candidate value θ* that is close to θ^(s) (we will get to how to generate the candidate value in a minute). Should we set θ^(s+1) = θ* or not?
- Well, we should probably compute $\pi(\theta^*|y)$ and see if $\pi(\theta^*|y) > \pi(\theta^{(s)}|y)$. Equivalently, look at $r = \frac{\pi(\theta^*|y)}{\pi(\theta^{(s)}|y)}$.
- By the way, notice that

$$egin{aligned} r &= rac{\pi(heta^\star|y)}{\pi(heta^{(s)}|y)} = rac{p(y| heta^\star)\pi(heta^\star)}{p(y)} \div rac{p(y| heta^{(s)})\pi(heta^{(s)})}{p(y)} \ &= rac{p(y| heta^\star)\pi(heta^\star)}{p(y)} imes rac{p(y)}{p(y| heta^{(s)})\pi(heta^{(s)})} = rac{p(y| heta^\star)\pi(heta^\star)}{p(y| heta^{(s)})\pi(heta^{(s)})}, \end{aligned}$$

which does not depend on the marginal likelihood we don't know!

- $\bullet \ \operatorname{lf} r > 1$
 - Intuition: θ^(s) is already a part of the density we desire and the density at θ^{*} is even higher than the density at θ^(s).
 - Action: set $\theta^{(s+1)} = \theta^{\star}$
- If r < 1,
 - Intuition: relative frequency of values on our group \$\theta^{(1)}\$, \dots, \$\theta^{(s)}\$ equal to \$\theta^{*}\$ should be \$\approx r = \frac{\pi(\theta^{*}|y)}{\pi(\theta^{(s)}|y)}\$. For every \$\theta^{(s)}\$, include only a fraction of an instance of \$\theta^{*}\$.
 - Action: set $\theta^{(s+1)} = \theta^*$ with probability r and $\theta^{(s+1)} = \theta^{(s)}$ with probability 1 r.



- This is the basic intuition behind the Metropolis algorithm.
- Where should the proposed value θ^* come from?
- Sample θ^{*} close to the current value θ^(s) using a symmetric proposal distribution g[θ^{*}|θ^(s)]. g is actually a "family of proposal distributions", indexed by the specific value of θ^(s).
- Here, symmetric means that $g[heta^{\star}| heta^{(s)}] = g[heta^{(s)}| heta^{\star}].$
- The symmetric proposal is usually very simple with density concentrated near $\theta^{(s)}$, for example, $\mathcal{N}(\theta^{\star}; \theta^{(s)}, \delta^2)$ or $\text{Unif}(\theta^{\star}; \theta^{(s)} \delta, \theta^{(s)} + \delta)$.
- After obtaining θ^* , either add it or add a copy of $\theta^{(s)}$ to our current set of values, depending on the value of r.



- The algorithm proceeds as follows:
 - 1. Given $heta^{(1)},\ldots, heta^{(s)}$, generate a candidate value $heta^\star \sim g[heta^\star| heta^{(s)}].$
 - 2. Compute the acceptance ratio

$$r = rac{\pi(heta^\star|y)}{\pi(heta^{(s)}|y)} = rac{p(y| heta^\star)\pi(heta^\star)}{p(y| heta^{(s)})\pi(heta^{(s)})}$$

3. Set

$$heta^{(s+1)} = egin{cases} heta^\star & ext{ with probability } \min(r,1) \ heta^{(s)} & ext{ with probability } 1-\min(r,1) \end{cases}$$

which can be accomplished by sampling $u \sim U(0,1)$ independently and setting

$$heta^{(s+1)} = egin{cases} heta^{\star} & ext{if} \quad u < r \ heta^{(s)} & ext{if} \quad ext{otherwise} \end{cases}$$



- Once we obtain the samples, then we are back to using Monte Carlo approximations for quantities of interest.
- That is, we can again approximate posterior means, quantiles, and other quantities of interest using the empirical distribution of our sampled values.
- Some notes:
 - The Metropolis chain ALWAYS moves to the proposed θ^* at iteration s+1 if θ^* has higher target density than the current $\theta^{(s)}$.
 - Sometimes, it also moves to a θ^{*} value with lower density in proportion to the density value itself.
 - This leads to a random, Markov process than naturally explores the space according to the probability defined by π(θ|y), and hence generates a sequence that, while dependent, eventually represents draws from π(θ|y).



METROPOLIS ALGORITHM: CONVERGENCE

- We will not cover the convergence theory behind Metropolis chains in detail, but below are a few notes for those interested:
 - The Markov process generated under this condition is ergodic and has a limiting distribution.
 - Here, think of ergodicity as meaning that the chain can move anywhere at each step, which is ensured, for example, if g[\(\theta^{\pm(s)}\)] > 0 everywhere!
 - By construction, it turns out that the Metropolis chains are reversible, so that convergence to $\pi(\theta|y)$ is assured.
 - Think of reversibility as being equivalent to symmetry of the joint density of two consecutive θ^(s) and θ^(s+1) in the stationary process, which we do have by using a symmetric proposal distribution.
- If you want to learn more about convergence of MCMC chains, consider taking one of the courses on stochastic processes, or Markov chain theory.



METROPOLIS ALGORITHM: TUNING

- Correlation between samples can be adjusted by selecting optimal δ (i.e., spread of the distribution) in the proposal distribution
- Decreasing correlation increases the effective sample size, increasing rate of convergence, and improving the Monte Carlo approximation to the posterior.
- However,
 - δ too small leads to $r \approx 1$ for most proposed values, a high acceptance rate, but very small moves, leading to highly correlated chain.
 - δ too large can get "stuck" at the posterior mode(s) because θ* can get very far away from the mode, leading to a very low acceptance rate and again high correlation in the Markov chain.
- Thus, good to implement several short runs of the algorithm varying δ and settle on one that yields acceptance rate in the range of 25-50%.
- orcozur
 Burn-in (and thinning) is even more important here!

METROPOLIS IN ACTION

Back to our example with

$$\pi(heta|y) \propto \exp^{-rac{1}{2} heta^2} + rac{1}{2} \exp^{-rac{1}{2}(heta-3)^2}$$

 $\pi(\theta|y)$



Move to the R script here.



WHAT'S NEXT?

MOVE ON TO THE READINGS FOR THE NEXT MODULE!

