STA 360/602L: MODULE 8.2

FINITE MIXTURE MODELS: UNIVARIATE CATEGORICAL DATA

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MULTINOMIAL MODEL RECAP

• Suppose $y_i, \ldots, y_n | oldsymbol{ heta} \stackrel{iid}{\sim} ext{Categorical}(oldsymbol{ heta})$, then

$$\Pr[y_i=d|oldsymbol{ heta}]=\prod_{d=1}^D heta_d^{1[y_i=d]}$$

• With prior $\pi[{m heta}] = {
m Dirichlet}({m lpha})$, we have

$$\pi[oldsymbol{ heta}] \propto \prod_{d=1}^D heta_j^{lpha_j-1}, \hspace{0.3cm} lpha_j > 0 \hspace{0.3cm} ext{for all} \hspace{0.3cm} d=1,\ldots,D.$$

So that the posterior is

$$\pi(oldsymbol{ heta}|Y) = ext{Dirichlet}(lpha_1+n_1,\ldots,lpha_D+n_D)$$

- However, what if our data actually comes from K different subpopulations of groups of people?
- For example, if our data comes from men and women, and we don't expect marginal independence across the two groups (vote turnout, income, etc), then we have a mixture of distributions.

FINITE MIXTURE OF MULTINOMIALS

- With our data coming from a "combination" or "mixture" of subpopulations, we no longer have independence across all observations, so that the likelihood p[Y|θ] ≠ ∏_{i=1}ⁿ ∏_{d=1}^D θ_j^{1[y_i=d]}.
- However, we can still have "conditional independence" within each group.
- Unfortunately, we do not always know the indexes for those groups.
- That is, we know our data contains K different groups, but we actually do not know which observations belong to which groups.
- Solution: introduce a latent variable z_i representing the group/cluster indicator for each observation i, so that each $z_i \in \{1, \ldots, K\}$.
- This is a form of data augmentation, but we will define that properly later.



FINITE MIXTURE OF MULTINOMIALS

• Given the cluster indicator z_i for observation i, write

Then, the marginal probabilities we care about will be

$$egin{aligned} & heta_d = \Pr(y_i = d) \ &= \sum_{k=1}^K \Pr(y_i = d | z_i = k) \cdot \Pr(z_i = k) \ &= \sum_{k=1}^K \lambda_k \cdot \psi_{k,d}, \end{aligned}$$

which is a finite mixture of multinomials, with the weights given by λ_k .



- Write
 - $oldsymbol{\lambda} = (\lambda_1, \dots, \lambda_K)$, and
 - $\psi = \{\psi_{z_i,d}\}$ to be a $K \times D$ matrix of probabilities, where each kth row is the vector of probabilities for cluster k.
- The observed data likelihood is

$$p\left[Y=(y_1,\ldots,y_n)|Z=(z_1,\ldots,z_n),oldsymbol{\psi},oldsymbol{\lambda}
ight] = \prod_{i=1}^n \prod_{d=1}^D \Pr\left(y_i=d|z_i,\psi_{z_i,d}
ight) \ = \prod_{i=1}^n \prod_{d=1}^D \psi_{z_i,d}^{1[y_i=d|z_i]},$$

which includes products (and not the sums in the mixture pdf), and as you will see, makes sampling a bit easier.

• Next we need priors.



• First, for $\lambda = (\lambda_1, \dots, \lambda_K)$, the vector of cluster probabilities, we can use a Dirichlet prior. That is,

$$\pi[oldsymbol{\lambda}] = ext{Dirichlet}(lpha_1,\ldots,lpha_K) \propto \prod_{k=1}^K \lambda_k^{lpha_k-1}$$

• For ψ , we can assume independent Dirichlet priors for each cluster vector $\psi_k = (\psi_{k,1}, \dots, \psi_{k,D})$. That is, for each $k = 1, \dots, K$,

$$\pi[oldsymbol{\psi}_k] = ext{Dirichlet}(a_1,\ldots,a_d) \propto \prod_{d=1}^D \psi_{k,d}^{a_d-1}.$$

• Finally, from our distribution on the z_i 's, we have

$$p\left[Z=(z_1,\ldots,z_n)|oldsymbol{\lambda}
ight]=\prod_{i=1}^n\prod_{k=1}^K\lambda_k^{1[z_i=k]}.$$



- Note that the unobserved variables and parameters are Z = (z₁,..., z_n)
 , ψ, and λ.
- So, the joint posterior is

$$egin{aligned} &\pi\left(Z, oldsymbol{\psi}, oldsymbol{\lambda}
ight) &\propto p\left[Y|Z, oldsymbol{\psi}, oldsymbol{\lambda}
ight] \cdot p(Z|oldsymbol{\psi}, oldsymbol{\lambda}) & \propto \left(\prod_{i=1}^n \prod_{d=1}^D p\left(y_i = d|z_i, oldsymbol{\psi}_{z_i,d}
ight)
ight] \cdot p(Z|oldsymbol{\lambda}) \cdot \pi(oldsymbol{\psi}) \cdot \pi(oldsymbol{\lambda}) & & \propto \left(\prod_{i=1}^n \prod_{d=1}^D oldsymbol{\psi}_{z_i,d}^{1[y_i = d|z_i]}
ight) & & imes \left(\prod_{i=1}^n \prod_{k=1}^K \lambda_k^{1[z_i = k]}
ight) & & imes \left(\prod_{k=1}^K \prod_{d=1}^D oldsymbol{\psi}_{k,d}^{a_d - 1}
ight) & & imes \left(\prod_{k=1}^K \lambda_k^{lpha_k - 1}
ight) & & imes \left(\prod_{k=1}^K \lambda_k^{lpha_k - 1}
ight). \end{aligned}$$



- First, we need to sample the z_i's, one at a time, from their full conditionals.
- For i = 1, ..., n, sample $z_i \in \{1, ..., K\}$ from a categorical distribution (multinomial distribution with sample size one) with probabilities

$$egin{aligned} & \Pr[z_i = k | \boldsymbol{y}_i, \boldsymbol{\psi}_k, \lambda_k] \ &= rac{\Pr[y_i, z_i = k | \boldsymbol{\psi}_k, \lambda_k]}{\sum\limits_{l=1}^K \Pr[y_i, z_i = l | \boldsymbol{\psi}_l, \lambda_l]} \ &= rac{\Pr[y_i | z_i = k, \boldsymbol{\psi}_k] \cdot \Pr[z_i = k, \lambda_k]}{\sum\limits_{l=1}^K \Pr[y_i | z_i = l, \boldsymbol{\psi}_l] \cdot \Pr[z_i = l, \lambda_l]} \ &= rac{\boldsymbol{\psi}_{k,d} \cdot \lambda_k}{\sum\limits_{l=1}^K \boldsymbol{\psi}_{l,d} \cdot \lambda_l}. \end{aligned}$$



Next, sample each cluster vector $oldsymbol{\psi}_k = (\psi_{k,1}, \dots, \psi_{k,D})$ from

$$\begin{split} \pi[\boldsymbol{\psi}_{k}|\ldots] &\propto \pi\left(Z,\boldsymbol{\psi},\boldsymbol{\lambda}|Y\right) \\ &\propto \left(\prod_{i=1}^{n}\prod_{d=1}^{D}\psi_{z_{i},d}^{1[y_{i}=d|z_{i}]}\right) \cdot \left(\prod_{i=1}^{n}\prod_{k=1}^{K}\lambda_{k}^{1[z_{i}=k]}\right) \cdot \left(\prod_{k=1}^{K}\prod_{d=1}^{D}\psi_{k,d}^{a_{d}-1}\right) \\ &\propto \left(\prod_{d=1}^{D}\psi_{k,d}^{n_{k,d}}\right) \cdot \left(\prod_{d=1}^{D}\psi_{k,d}^{a_{d}-1}\right) \\ &= \left(\prod_{d=1}^{D}\psi_{k,d}^{a_{d}+n_{k,d}-1}\right) \\ &\equiv \text{Dirichlet}\left(a_{1}+n_{k,1},\ldots,a_{D}+n_{k,D}\right). \end{split}$$

where $n_{k,d} = \sum_{i:z_i=k} 1[y_i = d]$, the number of individuals in cluster k that are assigned to category d of the levels of y.



• Finally, sample $oldsymbol{\lambda} = (\lambda_1, \dots, \lambda_K)$, the vector of cluster probabilities from

$$\begin{split} \pi[\boldsymbol{\lambda}|\ldots] &\propto \pi\left(Z, \boldsymbol{\psi}, \boldsymbol{\lambda}|Y\right) \\ &\propto \left(\prod_{i=1}^{n} \prod_{d=1}^{D} \psi_{z_{i},d}^{1[y_{i}=d|z_{i}]}\right) \cdot \left(\prod_{i=1}^{n} \prod_{k=1}^{K} \lambda_{k}^{1[z_{i}=k]}\right) \cdot \left(\prod_{k=1}^{K} \prod_{d=1}^{D} \psi_{k,d}^{\alpha_{d}-1}\right) \cdot \left(\prod_{k=1}^{K} \lambda_{k}^{\alpha_{k}-1}\right) \\ &\propto \left(\prod_{i=1}^{n} \prod_{k=1}^{K} \lambda_{k}^{1[z_{i}=k]}\right) \cdot \left(\prod_{k=1}^{K} \lambda_{k}^{\alpha_{k}-1}\right) \\ &\propto \left(\prod_{k=1}^{K} \lambda_{k}^{n_{k}}\right) \cdot \left(\prod_{k=1}^{K} \lambda_{k}^{\alpha_{k}-1}\right) \\ &\propto \left(\prod_{k=1}^{K} \lambda_{k}^{\alpha_{k}+n_{k}-1}\right) \\ &\equiv \text{Dirichlet}\left(\alpha_{1}+n_{1},\ldots,\alpha_{K}+n_{K}\right), \end{split}$$
where $n_{k} = \sum_{i=1}^{n} 1[z_{i}=k]$, the number of individuals assigned to cluster k .

WHAT'S NEXT?

MOVE ON TO THE READINGS FOR THE NEXT MODULE!

